

● 协方差和相关系数的性质

协方差的性质

□ $\text{cov}(X, Y) = \text{cov}(Y, X) = E(XY) - E(X)E(Y)$

证
$$\begin{aligned} \text{cov}(X, Y) &= E[(X - EX)(Y - EY)] \\ &= E[XY - X \cdot E(Y) - Y \cdot E(X) + E(X) \cdot E(Y)] \\ &= E(XY) - E(X) \cdot E(Y) - E(Y) \cdot E(X) + E(X) \cdot E(Y) \\ &= E(XY) - E(X)E(Y) \\ &= \text{cov}(Y, X) \end{aligned}$$

□ $\text{cov}(aX, bY) = ab \text{cov}(X, Y)$

证 $\text{cov}(aX, bY) = E[(aX - E(aX))(bY - E(bY))]$

$$= E[a(X - E(X)) \cdot b(Y - E(Y))]$$
$$= abE[(X - E(X)) \cdot (Y - E(Y))]$$
$$= ab \text{cov}(X, Y)$$

□ $\text{cov}(X+Y, Z) = \text{cov}(X, Z) + \text{cov}(Y, Z)$

证
$$\begin{aligned}\text{cov}(X+Y, Z) &= E\left\{\left[(X+Y) - E(X+Y)\right](Z - EZ)\right\} \\ &= E\left\{\left[(X - EX) + (Y - EY)\right](Z - EZ)\right\} \\ &= E\left[(X - EX)(Z - EZ) + (Y - EY)(Z - EZ)\right] \\ &= E\left[(X - EX)(Z - EZ)\right] + E\left[(Y - EY)(Z - EZ)\right] \\ &= \text{cov}(X, Z) + \text{cov}(Y, Z)\end{aligned}$$

□ $\text{cov}(X, X) = D(X)$

证 $\text{cov}(X, X) = E[(X - EX)(X - EX)]$
 $= D(X)$

□ $|\text{cov}(X, Y)|^2 \leq D(X)D(Y)$

当 $D(X) > 0, D(Y) > 0$ 时，当且仅当

$$P(Y - E(Y) = t_0(X - E(X))) = 1$$

时，等式成立

—Cauchy-Schwarz不等式

证 5 令

$$\begin{aligned}g(t) &= E[(X - E(X)) \cdot t - (Y - E(Y))]^2 \\&= E[(X - E(X))^2 \cdot t^2 - 2(X - E(X))(Y - E(Y)) \cdot t + (Y - E(Y))^2] \\&= D(X) \cdot t^2 - 2 \operatorname{cov}(X, Y) \cdot t + D(Y)\end{aligned}$$

对任何实数 t ,

$$g(t) \geq 0 \quad \longrightarrow$$

$$4 \operatorname{cov}^2(X, Y) - 4D(X)D(Y) \leq 0$$

即 $|\operatorname{cov}(X, Y)|^2 \leq D(X)D(Y)$

等号成立 \longleftrightarrow $g(t) = 0$ 有两个相等的实零点

等号成立, 即 $|\text{cov}(X, Y)|^2 = D(X)D(Y)$

此时, 零点为

$$t_0 = -\frac{-2 \text{cov}(X, Y)}{2D(X)} = \frac{\text{cov}(X, Y)}{D(X)}$$

$$= \frac{\sqrt{D(X) \cdot D(Y)}}{D(X)} \left(\text{或} \frac{-\sqrt{D(X) \cdot D(Y)}}{D(X)} \right)$$

$$= \sqrt{\frac{D(Y)}{D(X)}} \left(\text{或} -\sqrt{\frac{D(Y)}{D(X)}} \right)$$

此时,

$$g(t_0) = 0 \text{ 即}$$

$$E[(Y - E(Y)) - t_0(X - E(X))]^2 = 0 \quad \}$$

可以证明 $E[(Y - E(Y)) - t_0(X - E(X))] = 0$

$$\longleftrightarrow D[(Y - E(Y)) - t_0(X - E(X))] = 0$$

$$\longleftrightarrow P[(Y - E(Y)) - t_0(X - E(X))] = 0] = 1$$

$$P[(Y - E(Y)) - t_0(X - E(X)) = 0] = 1$$

即

$$P[(Y - E(Y)) = t_0(X - E(X))] = 1$$

即 Y 与 X 有线性关系的概率等于 1，这种线性关系为

$$P[(Y - E(Y)) = \pm \sqrt{\frac{D(Y)}{D(X)}}(X - E(X))] = 1$$

$$P\left(\frac{Y - E(Y)}{\sqrt{D(Y)}} = \pm \frac{X - E(X)}{\sqrt{D(X)}}\right) = 1$$

标准化随机变量

设随机变量 X 的期望 $E(X)$ 、方差 $D(X)$ 都存在，且 $D(X) \neq 0$ ，则称

$$X^* = \frac{X - E(X)}{\sqrt{D(X)}}$$

为 X 的标准化随机变量。显然，

$$E(X^*) = 0, \quad D(X^*) = 1$$

$$P(Y^* = \pm X^*) = 1$$

完全类似地可以证明

$$E^2(XY) \leq E(X^2)E(Y^2)$$

当 $E(X^2) > 0, E(Y^2) > 0$ 时，当且仅当

$$P(Y = t_0 X) = 1$$

时，等式成立

$$g(t) = E[Y - tX]^2$$

已知 $E(X^2) = 0$

求证 $E(X) = 0$

证明：

由 $D(X) = E(X^2) - E^2(X) \geq 0$

知 $E^2(X) \leq E(X^2)$

故当 $E(X^2) = 0$ 知 $E(X) = 0$

相关系数的性质

□ $|\rho_{XY}| \leq 1$

证明：由 $|\text{cov}(X, Y)|^2 \leq D(X)D(Y)$

及 $\rho_{XY} = \frac{\text{cov}(X, Y)}{\sqrt{D(X)} \cdot \sqrt{D(Y)}}$

知 $\rho_{XY}^2 = \frac{|\text{cov}(X, Y)|^2}{D(X)D(Y)} \leq 1$

故 $|\rho_{XY}| \leq 1$

□ $|\rho_{XY}|=1 \iff$ Cauchy-Schwarz不等式的等号成立

\iff 即 Y 与 X 有线性关系的概率等于1，这种线性关系为

$$P\left(\frac{Y - E(Y)}{\sqrt{D(Y)}} = \pm \frac{X - E(X)}{\sqrt{D(X)}}\right) = 1$$

$$\rho_{XY} = 1 \quad \longrightarrow$$

$$t_0 = - \frac{-2 \operatorname{cov}(X, Y)}{2D(X)} = \frac{\operatorname{cov}(X, Y)}{D(X)}$$

$$= \frac{\rho \cdot \sqrt{D(X) \cdot D(Y)}}{D(X)} = \sqrt{\frac{D(Y)}{D(X)}}$$

$$P[(Y - E(Y)) = \sqrt{\frac{D(Y)}{D(X)}}(X - E(X))] = 1$$

$$P\left(\frac{Y - E(Y)}{\sqrt{D(Y)}} = \frac{X - E(X)}{\sqrt{D(X)}}\right) = 1$$

$$P(Y^* = X^*) = 1$$

$$\rho_{XY} = -1 \longrightarrow$$

$$t_0 = -\frac{-2 \operatorname{cov}(X, Y)}{2D(X)} = \frac{\operatorname{cov}(X, Y)}{D(X)}$$

$$= \frac{\rho \cdot \sqrt{D(X) \cdot D(Y)}}{D(X)} = -\sqrt{\frac{D(Y)}{D(X)}}$$

$$P[(Y - E(Y)) = -\sqrt{\frac{D(Y)}{D(X)}}(X - E(X))] = 1$$

$$P\left(\frac{Y - E(Y)}{\sqrt{D(Y)}} = -\frac{X - E(X)}{\sqrt{D(X)}}\right) = 1$$

$$P(Y^* = -X^*) = 1$$

□ $\rho_{XY} = 0 \iff X, Y$ 不相关

$\iff \text{cov}(X, Y) = 0$

$\iff E(XY) = E(X)E(Y)$

$\iff D(X \pm Y) = D(X) + D(Y)$

X, Y 相关时,

$$D(X + Y) = D(X) + D(Y) + 2 \text{cov}(X, Y)$$

$$D(X - Y) = D(X) + D(Y) - 2 \text{cov}(X, Y)$$

X, Y 相互独立 $\xrightarrow{\quad}$ X, Y 不相关
 $\xleftarrow{\quad}$

若 X, Y 服从二维正态分布,

X, Y 相互独立 $\iff X, Y$ 不相关

例5 设 $(X, Y) \sim N(1, 4; 1, 4; 0.5)$,
 $Z = X + Y$, 求 ρ_{XZ}

解 $E(X) = E(Y) = 1$, $\rho_{XY} = \frac{1}{2}$,
 $D(X) = D(Y) = 4$,

$$\text{cov}(X, Y) = \rho_{XY} \cdot \sqrt{D(X)} \cdot \sqrt{D(Y)} = 2$$

$$\begin{aligned}\text{cov}(X, Z) &= \text{cov}(X, X) + \text{cov}(X, Y) \\ &= D(X) + \rho_{XY} \cdot \sqrt{D(X)} \cdot \sqrt{D(Y)}\end{aligned}$$

$$= 4 + \frac{1}{2} \cdot \sqrt{4} \cdot \sqrt{4} = 6$$

$$\begin{aligned}
D(Z) &= D(X + Y) \\
&= D(X) + D(Y) + 2 \operatorname{cov}(X, Y) \\
&= D(X) + D(Y) + 2 \cdot \rho_{XY} \cdot \sqrt{D(X)} \cdot \sqrt{D(Y)} \\
&= 4 + 4 + 2 \cdot \frac{1}{2} \sqrt{4} * \sqrt{4} = 12 \\
\rho_{XZ} &= \frac{\operatorname{cov}(X, Z)}{\sqrt{D(X)} \cdot \sqrt{D(Z)}} \\
&= \frac{6}{2 \cdot \sqrt{12}} \\
&= \frac{\sqrt{3}}{2}
\end{aligned}$$

例6 设随机变量 X 的概率密度函数为

$$f(x) = \frac{1}{2}e^{-|x|}, -\infty < x < +\infty$$

- (1) $E(|X|), D(|X|)$
- (2) 求 $\text{cov}(X, |X|)$, 问 X 与 $|X|$ 是否不相关.
- (3) 问 X 与 $|X|$ 是否独立? 为什么?

解 (1)
$$E(|X|) = \int_{-\infty}^{+\infty} |x| \frac{1}{2} e^{-|x|} dx$$

$$= 2 \cdot \int_0^{+\infty} x \frac{1}{2} e^{-x} dx$$

$$= 1$$

$$\begin{aligned}
E(|X|^2) &= \int_{-\infty}^{+\infty} |x|^2 \frac{1}{2} e^{-|x|} dx \\
&= 2 \cdot \int_0^{+\infty} x^2 \frac{1}{2} e^{-x} dx &= \int_0^{+\infty} x^2 e^{-x} dx \\
&= 2
\end{aligned}$$

$$D(|X|) = E(|X|^2) - E^2(|X|)$$

$$=2-1^2=1$$

$$\begin{aligned}
(2) \quad E(X \cdot | X|) &= \int_{-\infty}^{+\infty} x \cdot |x| \frac{1}{2} e^{-|x|} dx \\
&= \frac{1}{2} \int_{-\infty}^0 -x^2 e^x dx + \frac{1}{2} \int_0^{+\infty} x^2 e^{-x} dx \\
&= \frac{1}{2} \int_{+\infty}^0 -y^2 e^{-y} d(-y) + \frac{1}{2} \int_0^{+\infty} x^2 e^{-x} dx \\
&= -\frac{1}{2} \int_0^{+\infty} y^2 e^{-y} dy + \frac{1}{2} \int_0^{+\infty} x^2 e^{-x} dx \\
&= 0
\end{aligned}$$

$$\begin{aligned}
E(X) &= \int_{-\infty}^{+\infty} x \frac{1}{2} e^{-|x|} dx \\
&= \frac{1}{2} \int_{-\infty}^0 x e^x dx + \frac{1}{2} \int_0^{+\infty} x e^{-x} dx \\
&= 0
\end{aligned}$$

已证: $E(|X|) = 1$

$$\implies \text{cov}(X, |X|) = E(X|X|) - E(X)E(|X|) = 0$$

X 与 $|X|$ 不相关.

$$(3) \quad P(X < -2, |X| < 1) = 0$$

$$\begin{aligned} P(X < -2) &= \int_{-\infty}^{-2} f(x) dx = \int_{-\infty}^{-2} \frac{1}{2} e^{-|x|} dx \\ &= \frac{1}{2} \int_{-\infty}^{-2} e^x dx = \frac{1}{2} e^{-2} \end{aligned}$$

$$\begin{aligned} P(|X| < 1) &= \int_{-1}^1 f(x) dx = \int_{-1}^1 \frac{1}{2} e^{-|x|} dx \\ &= 2 \cdot \frac{1}{2} \int_0^1 e^{-x} dx = 1 - e^{-1} \end{aligned}$$

显然 $P(X < -2, |X| < 1) \neq P(X < -2)P(|X| < 1)$

因而 X 与 $|X|$ 不独立